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# 12 Count-Min Sketch and Apriori Algorithm (and Bloom Filters)

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Many streaming algorithms use random hashing functions to compress data. They basically randomly map some data items on top of each other. This leads to some error, but if one is careful, the large important items show through. The unimportant items get distributed randomly, and essentially add a layer of noise. Because it is random, this noise can be bounded.

An added benefit of the use of hash functions is the obliviousness. The algorithm is completely independent of the order or content of the data. This means its easy change items later if there was a mistake, or sketch different parts of data separately, and then combine them together later. Or if data is reordered due to race conditions, it has no effect on the output.

Next we also consider a variant of the frequent items problem: the frequent itemset problem, where we seek to find items that commonly occur together. Again, we will see instance where important and common itemsets show through above the noise.

This fits into the classic problem of *association rule mining*. The basic problem is posed as follows: We have a large set of  $m$  tuples  $\{T_1, T_2, \dots, T_m\}$ , each tuple  $T_j = \{t_{j,1}, t_{j,2}, \dots, t_{j,k}\}$  has a small number (not all the same  $k$ ) of items from a domain  $[n]$ . Think of the items as products in a grocery store and the tuples as things bought in the same purchase. Then the goal is to find times which frequently co-occur together. Think of what items to people frequently buy together: a famous example is “diapers” and “beer” found in real data.

Again, like finding frequent items (heavy hitters), we assume  $m$  and  $n$  is very large (maybe not quite so large as before), but we don’t want to check all pairs  $\binom{n}{2} \approx n^2/2$  against all  $m$  tuples since that would take roughly  $n^2m/2$  time and be way too long. Moreover,  $m$  does not fit in memory, so we only want to scan over it, in one pass (or a small number of passes).

## 12.1 Count-Min Sketch

In contrast to the Misra-Gries algorithm covered last lecture, we describe a completely different way to solve the HEAVY-HITTER problem. It is called the Count-Min Sketch [Cormode + Muthukrishnan 2005] [5].

Start with  $t$  independent (random) hash functions  $\{h_1, \dots, h_t\}$  where each  $h_h : [n] \rightarrow [k]$ .

Now we store a 2d array of counters for  $t = \log(1/\delta)$  and  $k = 2/\epsilon$ :

$h_1$	$C_{1,1}$	$C_{1,2}$	$\dots$	$C_{1,k}$
$h_2$	$C_{2,1}$	$C_{2,2}$	$\dots$	$C_{2,k}$
$\dots$	$\dots$	$\dots$	$\dots$	$\dots$
$h_t$	$C_{t,1}$	$C_{t,2}$	$\dots$	$C_{t,k}$

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**Algorithm 12.1.1** Count-Min( $A$ )

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Set all  $C_{i,j} = 0$

**for**  $i = 1$  **to**  $m$  **do**

**for**  $j = 1$  **to**  $t$  **do**

$C_{j,h_j(a_i)} = C_{j,h_j(a_i)} + 1$

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After running Algorithm 12.1.1 on a stream  $A$ , then on a query  $q \in [n]$  we can return

$$\hat{f}_q = \min_{j \in [t]} C_{j, h_j(q)}.$$

This is why it is called a *count-min sketch*.

**Analysis:** Clearly  $f_q \leq \hat{f}_q$  since each counter has everything for  $q$ , but may also have other stuff (on hash collisions).

Next we claim that  $\hat{f}_q \leq f_q + W$  for some over count value  $W$ . So how large is  $W$ ?

Consider just one hash function  $h_i$ . It adds to  $W$  when there is a collision  $h_i(q) = h_i(j)$ . This happens with probability  $1/k$ .

So we can create a random variable  $Y_{i,j}$  that represents the overcount caused on  $h_i$  for  $q$  because of element  $j \in [n]$ . That is, for each instance of  $j$ , it increments  $W$  by 1 with probability  $1/k$ , and 0 otherwise. Each instance of  $j$  has the same value  $h_i(j)$ , so we need to sum up all these counts. Thus

- $Y_{i,j} = \begin{cases} f_j & \text{with probability } 1/k \\ 0 & \text{otherwise.} \end{cases}$
- $\mathbf{E}[Y_{i,j}] = f_j/k$ .

Then let  $X_i$  be another random variable defined

- $X_i = \sum_{j \in [n], j \neq q} Y_{i,j}$ , and
- $\mathbf{E}[X_i] = \mathbf{E}[\sum_{j \neq q} Y_{i,j}] = \sum_{j \neq q} f_j/k = F_1/k = \varepsilon F_1/2$ .

Now we recall the Markov Inequality. For a random variable  $X$  and a value  $\alpha > 0$ , then  $\mathbf{Pr}[|X| \geq \alpha] \leq \mathbf{E}[|X|]/\alpha$ . Since  $X_i > 0$ , then  $|X_i| = X_i$ , and set  $\alpha = \varepsilon F_1$ . And note  $\mathbf{E}[|X|]/\alpha = (\varepsilon F_1/2)/(\varepsilon F_1) = 1/2$ . It follows that

$$\mathbf{Pr}[X_i \geq \varepsilon F_1] \leq 1/2.$$

But this was for just one hash function  $h_i$ . Now we extend this to  $t$  independent hash functions:

$$\begin{aligned} \mathbf{Pr}[\hat{f}_q - f_q \geq \varepsilon F_1] &= \mathbf{Pr}[\min_i X_i \geq \varepsilon F_1] = \mathbf{Pr}[\forall_{i \in [t]} (X_i \geq \varepsilon F_1)] \\ &= \prod_{i \in [t]} \mathbf{Pr}[X_i \geq \varepsilon F_1] \leq 1/2^t = \delta, \end{aligned}$$

since  $t = \log(1/\delta)$ .

So that gives us a PAC bound. The Count-Min Sketch for any  $q$  has

$$f_q \leq \hat{f}_q \leq f_q + \varepsilon F_1$$

where the first inequality always holds, and the second holds with probability at least  $1 - \delta$ .

**Space.** Since there are  $kt$  counters, and each require  $\log m$  space, then the total counter space is  $kt \log m$ . But we also need to store  $t$  hash functions, these can be made to take  $\log n$  space each. Then since  $t = \log(1/\delta)$  and  $k = 2/\varepsilon$  it follows the overall total space is  $t(k \log m + \log n) = ((2/\varepsilon) \log m + \log n) \log(1/\delta)$ .

**Turnstile Model:** There is a variation of streaming algorithms where each element  $a_i \in A$  can either add one or subtract one from corpus (like a turnstile at the entrance of a football game), but each count must remain positive. This Count-Min has the same guarantees in the turnstile model, but Misra-Gries does not.

## 12.2 Count Sketch

A predecessor of the Count-Min Sketch is the so-called Count Sketch [4]. Its structure is very similar to the Count-Min Sketch, it again maintains a 2d array of counters, but now with for  $t = \log(2/\delta)$  and  $k = 4/\varepsilon^2$ :

$h_1$	$C_{1,1}$	$C_{1,2}$	$\dots$	$C_{1,k}$
$h_2$	$C_{2,1}$	$C_{2,2}$	$\dots$	$C_{2,k}$
$\dots$	$\dots$	$\dots$	$\dots$	$\dots$
$h_t$	$C_{t,1}$	$C_{t,2}$	$\dots$	$C_{t,k}$

In addition to the  $t$  hash functions  $h_j : [n] \rightarrow [k]$  it maintains  $t$  sign hash functions  $s_j : [n] \rightarrow \{-1, +1\}$ . Then each hashed-to counter is incremented by  $s_j(a_i)$ . So it might add 1 or subtract 1.

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### Algorithm 12.2.1 Count-Min Sketch( $A$ )

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Set all  $C_{i,j} = 0$

**for**  $i = 1$  **to**  $m$  **do**

**for**  $j = 1$  **to**  $t$  **do**

$C_{j,h_j(a_i)} = C_{j,h_j(a_i)} + s_j(a_i)$

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To query this sketch, it takes the median of all values, instead of the minimum.

$$\hat{f}_q = \text{median}_{j \in [t]} \{C_{j,h_j(q)} \cdot s_j(q)\}.$$

Unlike the biased Count-Min Sketch, the other items hashed to the same counter as the query are unbiased. Half the time the values are added, and half the time they are subtracted. So then the median of all rows provides a better estimate. This insures the following bound with probability at least  $1 - \delta$  for all  $q \in [n]$ :

$$|f_q - \hat{f}_q| \leq \varepsilon F_2.$$

Note this required  $k = O(1/\varepsilon^2)$  instead of  $O(1/\varepsilon)$ , but usually the bound based on  $F_2 = \sqrt{\sum_j f_j^2}$  is much smaller than  $F_1 = \sum_j f_j$ , especially for skewed distributions. We will discuss so-called *heavy-tailed* distributions later in the class.

## 12.3 A-Priori Algorithm

We now describe the *A-Priori Algorithm* for finding frequent item sets, by Agrawal + Srikant [1]. The key idea is that any itemset that occurs frequently together must have each item (or any subset) occur at least as frequently.

**First Pass.** We first make one pass on all tuples, and keep a count for all  $n$  items. A hash table can be used. We set a threshold  $\varepsilon$  and only keep items that occur at least  $\varepsilon m$  times (that is in at least  $\varepsilon$  percent of the tuples). For any frequent itemset that occurs in at least  $100\varepsilon\%$  of the tuples, must have each item also occur in at least  $100\varepsilon\%$  of the tuples.

A reasonable choice of  $\varepsilon$  might be 0.01, so we only care about itemsets that occur in 1% of the tuples. Consider that there are only  $n_1$  items above this threshold. For instance if the maximum tuple size is  $k_{\max}$  then we know  $n_1 \leq k_{\max}/\varepsilon$ . For  $k_{\max} = 80$ , and  $\varepsilon = 0.01$ , then  $n_1 \leq 8000$ , easily small enough to fit in memory. Note this is independent of the  $n$  or  $m$ .

**Second Pass.** We now make a second pass over all tuples. On this pass we search for frequent pairs of items, specifically, those items which occur in at least an  $\varepsilon$ -fraction of all tuples. Both items must have been found in the first pass. So we need to consider only  $\binom{n_1}{2} \approx n_1^2/2$  pairs of counters for these pairs of elements.

After the pass, again we can then discard all pairs which occur less than an  $\varepsilon$ -fraction of all tuples. This remaining set is likely far less than  $n_1^2/2$ .

These remaining pairs are already quite interesting! They record all pairs that co-occur in more than an  $\varepsilon$ -fraction of purchases, and of course include those pairs which occur together even more frequently.

**Further Passes.** On the  $i$ th pass we can find sets of  $i$  items that occur together frequently (above an  $\varepsilon$ -threshold). For instance, on the third pass we only need to consider triples where all sub-pairs occur at least an  $\varepsilon$ -fraction of times themselves. These triples can be found as follows:

Sort all pairs  $(p, q)$  by their smaller indexed item (w.l.o.g. let this be  $p$ ). Then for each smaller indexed item  $p$ , consider all completions of this pair  $q$  (e.g. a triple  $(p, q, r)$ ). We only need to consider triples with  $(p, q, r)$  where  $p < q < r$ . Now for each pair  $(q, r)$ , check if the pair  $(p, r)$  also remains. Only triples  $(p, q, r)$  which pass all of these tests are given counters in the third pass.

This can be generalized to checking only  $k$  conditions in the  $k$ th pass, and the remaining triples form a lattice.

### 12.3.1 Example

Consider the following dataset where I want to find all itemsets that occur in at least 1/3 of all tuples (at least 4 times):

$$T_1 = \{1, 2, 3, 4, 5\}$$

$$T_2 = \{2, 6, 7, 9\}$$

$$T_3 = \{1, 3, 5, 6\}$$

$$T_4 = \{2, 6, 9\}$$

$$T_5 = \{7, 8\}$$

$$T_6 = \{1, 2, 6\}$$

$$T_7 = \{0, 3, 5, 6\}$$

$$T_8 = \{0, 2, 4\}$$

$$T_9 = \{2, 4\}$$

$$T_{10} = \{6, 7, 9\}$$

$$T_{11} = \{3, 6, 9\}$$

$$T_{12} = \{6, 7, 8\}$$

After the first pass I have the following counters:

0	1	<b>2</b>	<b>3</b>	4	5	<b>6</b>	<b>7</b>	8	<b>9</b>
2	3	<b>5</b>	<b>4</b>	3	3	<b>8</b>	<b>4</b>	2	<b>4</b>

So only  $n_1 = 5$  items survive  $\{2, 3, 6, 7, 9\}$ .

In pass 2 we consider  $\binom{n_1}{2} = 10$  pairs:  $\{(2, 3), (2, 6), (2, 7), (2, 9), (3, 6), (3, 7), (3, 9), (6, 7), (6, 9), (7, 9)\}$ . And we find the following counts:

(2, 3)	(2, 6)	(2, 7)	(2, 9)	(3, 6)	(3, 7)	(3, 9)	(6, 7)	<b>(6, 9)</b>	(7, 9)
1	3	1	2	3	0	1	3	<b>4</b>	2

We find that the only itemset pair that occurs in at least 1/3 of all baskets is (6, 9).

Thus there can be no itemset triple (or larger grouping) which occurs in all 1/3 of all tuples since then all of its pairs would need to be in 1/3 of all baskets, but there is only one such pair that satisfies that property.

We can now examine the association rules. And see that the count of item 6 is quite large 8, and is much bigger than that of item 9 which is only 4. Since there are 4 pairs (6, 9), then every time 9 occurs, 6 also occurs.

**Improvements.** We can do better than keeping a counter for each item. If we know that  $n_1$  will be at most a certain value (based on  $k_{\max}$  and  $\epsilon$ ), then we only need that many counters, and we can use Misra-Gries to find all of these items. Although we may also find some other items, so its advised to use maybe  $3n_1$  counters in Misra-Gries.

Other techniques have been developed to improve on this using Bloom Filters.

## 12.4 Bloom Filters

The goal of a Bloom Filter [2] is to maintain a subset  $S \subset [n]$  in small space. It must maintain all items, but may allow some items to appear in the set even if they are not. That is, it allows *false positives*, but not *false negatives*.

It maintains an array  $B$  of  $m$  bits. Each is initialized to 0. It uses  $k$  (random) hash functions  $\{h_1, h_2, \dots, h_k\} \in \mathcal{H}$ . It then runs streaming Algorithm 12.4.1 on  $S$ .

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**Algorithm 12.4.1** Bloom( $S$ )

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for  $x \in S$  do
  for  $j = 1$  to  $k$  do
    Set  $B[h_j(x)] = 1$ 
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Then on a query to see if  $y \in S$ , it returns YES only if for *all*  $j \in [k]$  that  $B[h_j(y)] = 1$ . Otherwise it returns NO.

So any item which was put in  $B$  it sets all associated hash values as 1, so on a query it will always return YES (no false negatives).

However, it may return YES for an item even if it does not appear in the set. It does not even need to have the exact same hash values from another item in the set, each hash collision could occur because a different item.

**Analysis:** We consider the case with  $|S| = n$  items, where we maintain  $m$  bits in  $B$ , and use  $k$  hash functions.

$$\begin{aligned} & 1 - 1/m \text{ Probability a bit not set to 1 by a single hash function} \\ & (1 - 1/m)^k \text{ Probability a bit not set to 1 by } k \text{ hash functions} \\ & (1 - 1/m)^{kn} \text{ On inserting } n \text{ elements, probability a bit is 0 } (\star) \\ & 1 - (1 - 1/m)^{kn} \text{ On inserting } n \text{ elements, probability a bit is 1} \\ & (1 - (1 - 1/m)^{kn})^k \text{ Probability a false positive (using } k \text{ hash functions)} \\ & \approx (1 - e^{-kn/m})^k \end{aligned}$$

The “right” value of  $k$  is about  $k \approx (m/n) \ln 2$ .

*(\*) This step is not quite right, although it is not too far off. It assumes independence of which bits are being set. Unfortunately, in most presentations of Bloom filters (and some peer-reviewed research papers) this is presented as the right analysis, but it is not. But it gives a pretty good illustration of how it works. The correct analysis is much more complicated [3].*

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# Bibliography

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